

The Economics of European Regions: Theory, Empirics, and Policy

Dipartimento di Economia e Management

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From Theory to Empirics in Solow model

From the standard Solow model we have that:

$$\dot{k} = sf(k, h) - (\delta + g_A + n)k, \quad (1)$$

where

$$k \equiv \frac{K}{AL}, \quad f \equiv F\left(\frac{K}{AL}, h\right) \quad \text{and} \quad f_k > 0, f_{kk} < 0 \quad (2)$$

and s is the exogenous saving/investment rate, h the level of human capital, δ the depreciation rate of physical capital, g_A the growth rate of technological change, and n the growth rate of employment.

Equilibrium level

Assuming a Cobb-Douglas function that $F(K, AH) = K^\alpha (AH)^{1-\alpha}$ then:

$$\dot{k} = sk^\alpha h^{1-\alpha} - (\delta + g_A + n)k; \quad (3)$$

$$k_t^{EQ} = h \left(\frac{s}{\delta + g_A + n} \right)^{1/(1-\alpha)}; \quad (4)$$

and:

$$y_t^{EQ} \equiv \frac{Y}{AL} = h \left(\frac{s}{\delta + g_A + n} \right)^{\alpha/(1-\alpha)}. \quad (5)$$

Log-linearization around the equilibrium

The generic one-sector growth model implies, to the first-order approximation, that:

$$\log(y_t) = (1 - e^{-\lambda t})\log(y^{EQ}) + e^{-\lambda t}\log(y_0), \quad (6)$$

where y_t is the level of GDP per worker in efficiency units at time t , and the parameter λ measures the **rate of convergence**. Eq. (6) is expressed in terms of the unobservable y_t .

Suppose that technological progress is evolving as:

$$A_t = A_0 e^{g_A t} \quad (7)$$

so that:

$$\log(A_t) = \log(A_0) + g_A t \quad (8)$$

Log-linearization around the equilibrium (2)

In order to describe the dynamics in terms of the observable GDP per worker, $\frac{Y_t}{L_t}$, we get:

$$\log\left(\frac{Y_t}{L_t}\right) - g_A t - \log(A_0) = (1 - e^{-\lambda t})\log(y^{EQ}) + e^{-\lambda t} \left[\log\left(\frac{Y_0}{L_0}\right) - \log(A_0) \right]. \quad (9)$$

Defining the average growth rate of GDP per worker, $\overline{g_{Y/L}}$, as:

$$\overline{g_{Y/L}} \approx \frac{\log\left(\frac{Y_t}{L_t}\right) - \log\left(\frac{Y_0}{L_0}\right)}{t} \quad (10)$$

and:

$$\beta \equiv -\frac{(1 - e^{-\lambda t})}{t} \quad (11)$$

Log-linearization around the euqilibrium (3)

Therefore we get:

$$\overline{g_{Y/L}} = g_A - \beta \log \left(\frac{Y_0}{L_0} \right) - \beta \log(y^{EQ}) - \beta \log(A_0) \quad (12)$$

Substituting the level of equilibrium y^{EQ} we get:

$$\begin{aligned} \overline{g_{Y/L}} &= g_A + \beta \log \left(\frac{Y_0}{L_0} \right) + \beta \frac{\alpha}{1-\alpha} \log(n + g_A + \delta) - \beta \frac{\alpha}{1-\alpha} \log(s) \\ &\quad - \beta \log(h) - \beta \log A_0 \end{aligned} \quad (13)$$

Issues for the estimation

- Initial condition on TFP not observable (A_0) \Rightarrow we need proxies!
- Constrained on parameters
- Intercept is an estimate of the growth rate of exogenous technological progress
- Conditional convergence occurs when $\beta < 0$ and $\beta > -1$ and depends on $t \Rightarrow$ speed of convergence
- All the determinants are exogenous and no relevant variable are omitted \Rightarrow OLS are unbiased estimators!

Cross-region estimation

$$\overline{g_{Y/L}}_i = \text{intercept} + \beta_0 \log(Y/L_{i,1991}) + \beta_1 \bar{s}_i + \beta_2 \bar{n}_i + \beta_3 \bar{h}_i + \epsilon_i \quad (14)$$

	Estimate	Std. Error	t-Stat.	P-value
(Intercept)	-0.0929	0.0123	-7.53	0.0000
β_0	-0.0154	0.0011	-14.57	0.0000
β_1	0.0027	0.0029	0.93	0.3532
β_2	-0.0146	0.0034	-4.31	0.0000
β_3	0.0204	0.0024	8.57	0.0000
Res.se=0.008956 (255) DF				
R-squared=0.6411, Adj.R-squared=0.6354				
F-stat.=112.6 (1,255) DF, p-value=< 2e ⁻¹⁶				

Endogeneity in cross-region regression

Simultaneity Problem

The fact that the right-hand-side variables are not exogenous, but are **jointly determined with the growth rate** (for example the level of investment is highly correlated with growth).

- *Estimation issue*: estimates can be biased.
- *Identification issue*: the value of β can fail to illustrate how initial conditions affect expected future income differences if the saving rate is itself function of income. Hence, $\beta \geq 0$ may be compatible with at least partial convergence, while $\beta < 0$ with economic divergence if physical and human capital accumulation for rich and poor are diverging across time.

Endogeneity in cross-region regression

Measurement Error

In this case we would like to measure the (partial) effect of a variable but we can **observe only an imperfect measure** \Rightarrow we introduce measurement error.

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Omitted Variables

Omitted variables appear when we would like to control for one or more additional variables but, usually because of data unavailability, we cannot include them in a regression model.

\Rightarrow one way to represent this situation is to write the regression equation considering the omitted variable as part of the error term.

Instrumental Variables and Two-Stage Least Squares

Consider the linear model:

$$y = \beta_0 + \beta_1 x_1 + \dots + \beta_K x_K + u \quad (15)$$

$$E(u) = 0, \text{Cov}(x_j, u) = 0, j = 1, 2, \dots, K - 1 \quad (16)$$

therefore x_K might be correlated with u . In other words, x_1, \dots, x_{K-1} are exogenous while x_K is potentially endogenous

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\Rightarrow OLS estimation generally results in **inconsistent** estimators of all the β_j if $\text{Cov}(x_K, u) \neq 0$

Instrumental Variables and Two-Stage Least Squares (2)

The method of instrumental variables (IV) provides a general solution to the problem of an endogenous explanatory variable. To use the IV approach with x_K endogenous, we need an observable variable, z_1 , not in equation (16) that satisfies two conditions:

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- z_1 must be uncorrelated with u : $Cov(z_1, u) = 0 \Rightarrow z_1$ is exogenous
- The second requirement involves the relationship between z_1 and the endogenous variable, x_K . Consider the regression of x_K on *all* the exogenous variables:

$$x_K = \delta_0 + \delta_1 x_1 + \delta_2 x_2 + \dots + \delta_{K-1} x_{K-1} + \theta_1 z_1 + e_K \quad (17)$$

where $E(e_K) = 0$ and e_K is uncorrelated with x_1, \dots, x_{K-1} and z_1
 $\Rightarrow \theta_1 \neq 0$

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where $E(e_K) = 0$ and e_K is uncorrelated with x_1, \dots, x_{K-1} and $z_1 \Rightarrow \theta_1 \neq 0$

z_1 is an **instrumental variable** candidate for x_K !

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This is called **first-stage regression**.

- 2 Run the OLS regression

$$y = \beta_0 + \beta_1 x_1 + \dots + \beta_K \hat{x}_K + u \quad (19)$$

This is called the **second-stage regression**, and it produces the $\hat{\beta}_j$

Control Function

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- CF uses extra regressors to break the correlation between endogenous explanatory variables and unobservables affecting the dependent variable.
- The method still relies on the availability of exogenous variables that do not appear in the structural equation

Control Function (2)

Consider the linear model:

$$y = \beta_0 + \beta_1 x_1 + \dots + \beta_K x_K + u \quad (20)$$

$$E(u) = 0, \text{Cov}(x_j, u) = 0, j = 1, 2, \dots, K - 1 \quad (21)$$

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Consider again the *reduced form* of x_K :

$$x_K = \delta_0 + \delta_1 x_1 + \delta_2 x_2 + \dots + \delta_{K-1} x_{K-1} + \theta_1 z_1 + e_K \quad (22)$$

with $\text{Cov}(x_j, e_K) = 0, j = 1, 2, \dots, K - 1$ and $\text{Cov}(z_1, e_K) = 0$.

Control Function (3)

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Write the linear projection of u on e_K as:

$$u = \rho e_K + \epsilon \quad (23)$$

By definition, $\text{Cov}(e_K, \epsilon) = 0$, $\text{Cov}(x_j, \epsilon) = 0$ and $\text{Cov}(z_1, \epsilon) = 0$ because u and e_K are both uncorrelated with x_j $j = 1, \dots, K_1$ and z_1 .

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Pluggin (9) in (6) we get:

$$y = \beta_0 + \beta_1 x_1 + \dots + \beta_K x_K + \rho e_K + \epsilon \quad (24)$$

where now e_K can be viewed as an explanatory variable in the equation, and $\text{Cov}(y, \epsilon) = 0$.

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\Rightarrow run OLS of y on x_j $j = 1, \dots, K_1$, z_1 and e_K using random sample.

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- given that we observe \mathbf{x}, z_1 we can estimate the model 25 by OLS
 \Rightarrow replace e_K with \hat{e}_K .

$$y = \beta_0 + \beta_1 x_1 + \dots + \beta_K x_K + \rho \hat{e}_K + \text{error} \quad (26)$$

where:

$$\text{error} = \epsilon + \rho(x_1, \dots, x_{K-1}, z_1) \left[(\hat{\delta}_0, \hat{\delta}_1, \dots, \hat{\delta}_{K-1}, \hat{\theta}) - (\delta_0, \delta_1, \dots, \delta_{K-1}, \theta) \right]$$

depends on the sampling error.

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\Rightarrow OLS estimator of (26) will be consistent!!

Control Function (5)

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- The OLS estimate of β_j and $j = 1, \dots, K$ are *identical* to the 2SLS.
- Test of endogeneity: $\rho = 0$.
- Problem: \hat{e}_K is a *generated regressor* \Rightarrow we need bootstrap for right standard errors!

Control Function: summarizing

- 1 Obtain the **residuals** \hat{e}_K from the regression:

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- 3 Test $\hat{\rho} = 0$.

References

- Wooldridge: Econometric Analysis of Cross Section and Panel Data; Chapter 5 and Chapter 6
- R file: endogeneity_EUregions.R in *EER_8Ott2018*.